

SOLUTIONS TO THE EXAM FOR RANDOM PROCESSES, 5P

January 18, 2003, 9.00 am – 13.00 pm

Max number of points: 30. **Bounds:** 12p \Rightarrow grade 3, 18p \Rightarrow grade 4, 24p \Rightarrow grade 5.

Allowed aids:

Sheet of formulae attached to the exam, calculator and Mathematics Handbook: Beta.

Examiner: Eric Järpe (035-16 76 53).

1. Show that if $G = \mathcal{F}(g)$ and $h(\tau) \equiv G(\tau)$ then är $H(f) \equiv g(-f)$
(where $H = \mathcal{F}(h)$ and $g = \mathcal{F}^{-1}(G)$). (3p)

Solution: Assume that $G = \mathcal{F}(g)$. Then

$$g(-\tau) = \mathcal{F}^{-1}(G)(-\tau) = \int_{\mathbb{R}} e^{i2\pi f \cdot (-\tau)} G(f) df = \int_{\mathbb{R}} e^{-i2\pi f \tau} G(f) df$$

i.e.

$$g(-f) = \int_{\mathbb{R}} e^{-i2\pi f \tau} G(\tau) d\tau = \int_{\mathbb{R}} e^{-i2\pi f \tau} h(\tau) d\tau = \mathcal{F}(h)(f) = H(f)$$

□

2. Show that if $\{X_t : t \in \mathbb{R}\}$ is weakly stationary and $Z_t = X_t$ för $t \in \{dk : k \in \mathbb{Z}\}$
then $S_Z(f) = \sum_{-\infty}^{\infty} S_X(f + \frac{k}{d})$ for $-\frac{1}{2d} < f < \frac{1}{2d}$. (4p)

Solution: (Se Lindgren-Rotzén, pp 75 – 76.)

3. Assume that $\{X_t\}$ och $\{Y_t\}$ are mutually independent Poisson processes with parameter $\lambda_X = 0.1$ and $\lambda_Y = 0.2$ respectively.

- (a) Calculate $P(X_{200} < Y_{101})$ approximatively. (3p)
(b) Show that the process $Z_t = X_{2t} - Y_t$ is not stationary. (3p)

Solution:

- (a) $X_t = \sum_{k=1}^t Z_k$ where Z_1, Z_2, \dots, Z_t are independent and $Z_k \in Po(0.1)$.
According to C.L.T. this means that approximately $X_t \in N(0.1t, \sqrt{0.1t})$.
 $Y_t = \sum_{k=1}^t W_k$ where W_1, W_2, \dots, W_t are independent and $W_k \in Po(0.2)$.
According to C.L.T. this means that approximately $Y_t \in N(0.2t, \sqrt{0.2t})$.
Further $E(X_{200} - Y_{101}) = E(X_{200}) - E(Y_{101}) = 0.1 \cdot 200 - 0.2 \cdot 101 = -0.2$
and, since $\{X_t\}$ and $\{Y_t\}$ are independent, we have that
 $V(X_{200} - Y_{101}) = V(X_{200}) + V(Y_{101}) = 0.1 \cdot 200 + 0.2 \cdot 101 = 40.2$
and hence $X_{200} - Y_{101} \in N(-0.2, \sqrt{40.2})$ so
 $P(X_{200} < Y_{101}) = P(X_{200} - Y_{101} < 0) = \Phi\left(\frac{0 - (-0.2)}{\sqrt{40.2}}\right) = \Phi(0.0315) = \underline{0.512}$

- (b) $Z_t = X_{2t} - Y_t$. For weak stationarity of $\{Z_t\}$ it is necessary that $E(Z_t) = m$ is not a function of t and that $C(Z_s, Z_t) = r_Z(s - t)$. If the process $\{Z_t\}$ is weakly stationary this does not imply strict stationarity *but* if it is *not* weakly stationary then it cannot be strictly stationary, i.e. then it can't be stationary in any sense. We have that $E(Z_t) = E(X_{2t} - Y_t) = E(X_{2t}) - E(Y_t) = 0.1 \cdot 2t - 0.2 \cdot t = 0$ (ok, it might still be weakly stationary...)

Since $\{X_t\}$ and $\{Y_t\}$ are Poisson processes,

$C(X_s, X_t) = 0.1 \min(s, t)$ and $C(Y_s, Y_t) = 0.2 \min(s, t)$. Therefore

$$\begin{aligned} C(Z_s, Z_t) &= C(X_{2s} - Y_s, X_{2t} - Y_t) = \\ &= C(X_{2s}, X_{2t}) - C(X_{2s}, Y_t) - C(Y_s, X_{2t}) + C(Y_s, Y_t) = \\ &= 0.1 \min(2s, 2t) - 0 - 0 + 0.2 \min(s, t) = 0.4 \min(s, t) \end{aligned}$$

However, $\min(s, t)$ is not a function of $s - t$ so $\{Z_t\}$ is not weakly stationary and thus $\{Z_t\}$ is neither weakly nor strictly stationary! \square

4. The spectral density of the process $\{X_t\}$ is $S_X(f) = |f|e^{-af^2}$. This signal will be sampled from with interval 0.3. What does a have to be for not risking that more than 10% of the signal is distorted? (3p)

Solution: Approximately 4 “tailparts” are distorted so

$$4 \int_{1/(2 \cdot 0.3)}^{\infty} R_X(f) df \leq 0.1 \int_{-\infty}^{\infty} R_X(f) df$$

$$\int_{1/0.6}^{\infty} |f|e^{-af^2} df = \frac{1}{2a} \int_{1/0.6}^{\infty} 2af e^{-af^2} df = \left\{ \begin{array}{ll} u = af^2 & f = 1/0.6 \\ du = 2af df & u = a(1/0.6)^2 = a/0.36 \end{array} \right\} = \\ = \frac{1}{2a} \int_{a/0.36}^{\infty} e^{-u} du = \frac{1}{2a} [-e^{-u}]_{a/0.36}^{\infty} = \frac{1}{2a} (0 - (-e^{-a/0.36})) = e^{-a/0.36} / (2a)$$

$$\int_{-\infty}^{\infty} |f|e^{-af^2} df = 2 \int_0^{\infty} f e^{-af^2} df = 2 \cdot \frac{1}{2a} \int_0^{\infty} e^{-u} du = \frac{1}{a} (0 - (-1)) = 1/a.$$

Therefore a must satisfy the condition

$$4 \cdot \frac{e^{-a/0.36}}{2a} \leq 0.1 \cdot \frac{1}{a}$$

$$\Rightarrow 2e^{-a/0.36} \leq 0.1$$

$$\Rightarrow -a/0.36 \leq \ln(0.1/2)$$

$$\Rightarrow a \geq -0.36 \ln 0.05 = 1.078 \quad (\text{Obs! Multiplying negative numbers switches the inequality.})$$

Answer: a must be larger than 1.078. \square

5. Assume that $\{X_t\}$ is an $AR(1)$ process with regression parameter $a_1 = -2/3$ and noise process $\{\epsilon_t\}$ where $\epsilon_t \in N(0, 1/2)$.

- (a) What does $E(X_t)$ and $V(X_t)$ have to be for the process $\{X_t\}$ to be strictly stationary? (3p)
- (b) Calculate expectation function and covariance function of the derivative process $\{X'_t\}$. (3p)

Solution: $\{X_t\}$ $AR(1)$ with $a_1 = -\frac{2}{3}$ means that $X_t - \frac{2}{3}X_{t-1} = \epsilon_t$ i.e. that $X_t = \frac{2}{3}X_{t-1} + \epsilon_t$ where $\epsilon_t \in N(0, \frac{1}{2})$.

- (a) By definition all AR -processes are weakly stationary. Since the noise is normally distributed the process is a normal process and therefore weak stationarity implies strict stationarity.

Weak stationarity implies that $E(X_t)$ and $V(X_t)$ are constant with respect to t . Since $X_t = \frac{2}{3}X_{t-1} + \epsilon_t$ we have, by taking the expectation of both sides of this equality, that

$$\begin{aligned} m &= E(X_t) = E\left(\frac{2}{3}X_{t-1} + \epsilon_t\right) = \frac{2}{3}E(X_t) + E(\epsilon_t) = \frac{2}{3}m + 0 \Rightarrow \\ &\Rightarrow m - \frac{2}{3}m = 0 \Rightarrow m = 0 \text{ and by taking the variance of both sides we} \\ &\text{have that } \sigma^2 = V(X_t) = V\left(\frac{2}{3}X_{t-1} + \epsilon_t\right) = \frac{4}{9}V(X_{t-1}) + V(\epsilon_t) = \frac{4}{9}\sigma^2 + \frac{1}{4} \Rightarrow \\ &\Rightarrow \sigma^2\left(1 - \frac{4}{9}\right) = \frac{1}{4} \Rightarrow \sigma^2 = \frac{1}{4} \cdot \frac{9}{5} = \frac{9}{20} = 0.45 \end{aligned}$$

Answer: $E(X_t) = 0$ and $V(X_t) = 0.45$

- (b) According to Thm (p 127, Lindgren-Rotzén) we have that if $\{X_t\}$ is weakly stationary then $m_{X'} = 0$ and $r_{X'}(\tau) = r''_X(\tau)$.

The expectation function is thus constantly 0.

Now we may calculate $R_X(\tau) = C(X_t, X_{t+\tau})$ and differentiate this twice and thereby get $R_{X'}(\tau)$ according to the theorem. We have that

$$\begin{aligned} C(X_t, X_{t+1}) &= C\left(X_t, \frac{2}{3}X_t + \epsilon_{t+1}\right) = \frac{2}{3}V(X_t) + 0 = \frac{2}{3} \cdot \frac{9}{20} \\ C(X_t, X_{t+2}) &= C\left(X_t, \frac{2}{3}X_{t+1} + \epsilon_{t+2}\right) = \frac{2}{3}C(X_t, X_{t+1}) = \frac{2}{3} \cdot \frac{2}{3} \cdot \frac{9}{20} = \left(\frac{2}{3}\right)^2 \frac{9}{20} \\ &\vdots \end{aligned}$$

$$C(X_t, X_{t+\tau}) = \left(\frac{2}{3}\right)^\tau \frac{9}{20} = r_X(\tau). \text{ Thus}$$

$$r'_{X'}(\tau) = D\left(e^{\tau \ln(2/3)} \frac{9}{20}\right) = \frac{9}{20} \ln \frac{2}{3} e^{\tau \ln(2/3)} = \frac{9}{20} \ln \frac{2}{3} \left(\frac{2}{3}\right)^\tau$$

$$r''_{X'}(\tau) = \frac{9}{20} \ln \frac{2}{3} \ln \frac{2}{3} \left(\frac{2}{3}\right)^\tau = \frac{9}{20} \left(\ln \frac{2}{3}\right)^2 \left(\frac{2}{3}\right)^\tau$$

Answer: $E(X'_t) = 0$ and $r_{X'}(\tau) = \frac{9}{20} \left(\ln \frac{2}{3}\right)^2 \left(\frac{2}{3}\right)^\tau$

□

6. Let the expectation function of $\{X_t\}$ be $m_X = 0$ and the covariance function be $R_X(\tau) = \max(0, \frac{1}{2} - |\tau|)$.

(a) Calculate the spectral density $S_X(f)$. (By all means, use the table!) (3p)

(b) The signal $\{Y_t\}$ is achieved by a linear transfer defined by $Y_t = \int_{t-1}^t X_u du$. Derive the spectral density of $\{Y_t\}$. (5p)

Solution:

(a) $m_X = 0$ and $R_X(\tau) = \max(0, \frac{1}{2} - |\tau|) = \begin{cases} \frac{1}{2} - |\tau| & \text{om } -\frac{1}{2} < \tau < \frac{1}{2} \\ 0 & \text{annars} \end{cases}$

According to the table we have that

$$g(\tau) = \begin{cases} 1 - \alpha|\tau| & \text{if } |\tau| \leq \frac{1}{\alpha} \\ 0 & \text{otherwise} \end{cases} \Rightarrow G(f) = \begin{cases} \frac{1}{\alpha} & \text{if } f = 0 \\ \frac{2\alpha}{(2\pi f)^2} (1 - \cos(\frac{2\pi f}{\alpha})) & \text{if } f \neq 0 \end{cases}$$

Now let $g(\tau) = \max(0, 1 - 2|\tau|)$ because then

$$G(f) = \begin{cases} \frac{1}{2} & \text{om } f = 0 \\ \frac{1 - \cos(\pi f)}{(\pi f)^2} & \text{om } f \neq 0 \end{cases}$$

Furthermore $\frac{1}{2}g(\tau) = \frac{1}{2}\max(0, 1 - 2|\tau|) = \max(0, \frac{1}{2} - |\tau|) = r_X(\tau)$

so since $G(f) = \int_{\mathbb{R}} e^{-i2\pi f\tau} g(\tau) d\tau$ the spectral density is

$$S_X(f) = \int_{\mathbb{R}} e^{-i2\pi f\tau} \frac{1}{2}g(\tau) d\tau = \frac{1}{2} \int_{\mathbb{R}} e^{-i2\pi f\tau} g(\tau) d\tau = \frac{1}{2}G(f) = \begin{cases} \frac{1}{4} & \text{if } f = 0 \\ \frac{1 - \cos(\pi f)}{2(\pi f)^2} & \text{if } f \neq 0 \end{cases}$$

(b) In order to derive the spectral density, S_Y , of the transferred signal we may derive the impulse response, h , and further the corresponding frequency function, H , and we will have the desired spectral density S_Y by the relation $S_Y = |H|^2 S_X$.

The impulse response is recognised by inspection of the linear transfer:

$Y_t = \int_{-\infty}^{\infty} h(u)X_{t-u} du$ so for the process $\{Y_t\}$ in this case we have that

$$Y_t = \int_{t-1}^t X_u du = \begin{cases} v = t - u & u = t \Rightarrow v = 0 \\ dv = -du & u = t - 1 \Rightarrow v = 1 \end{cases} = \int_1^0 X_{t-v} (-dv) = \int_0^1 X_{t-v} dv = \int_{-\infty}^{\infty} h(v)X_{t-v} dv \text{ if } h(v) = \begin{cases} 1 & \text{if } 0 < v < 1 \\ 0 & \text{otherwise} \end{cases}$$

Therefore the frequency function is

$$H(f) = \int_{-\infty}^{\infty} e^{-i2\pi f\tau} h(\tau) d\tau = \int_0^1 e^{-i2\pi f\tau} \cdot 1 d\tau = \left[\frac{e^{-i2\pi f\tau}}{-i2\pi f} \right]_0^1 = \frac{1 - e^{-i2\pi f}}{i2\pi f} \text{ if } f \neq 0.$$

$$\text{If } f = 0 \text{ then } H(f) = \int_0^1 e^0 \cdot 1 d\tau = 1 \text{ i.e. } H(f) = \begin{cases} \frac{1 - e^{-i2\pi f}}{i2\pi f} & \text{if } f \neq 0 \\ 1 & \text{if } f = 0 \end{cases}$$

$$S_Y(f) = |H(f)|^2 S_X(f) = \begin{cases} \left| \frac{1 - e^{-i2\pi f}}{i2\pi f} \right|^2 \frac{1 - \cos(\pi f)}{2(\pi f)^2} & \text{if } f \neq 0 \\ |1|^2 \cdot \frac{1}{4} & \text{if } f = 0 \end{cases}$$

Assume $f \neq 0$. Then with $e^{i\alpha} = \cos \alpha + i \sin \alpha$ we have that $S_Y(f) =$

$$= \left| \frac{1 - \cos(2\pi f) + i \sin(2\pi f)}{i2\pi f} \right|^2 \frac{1 - \cos(\pi f)}{2(\pi f)^2} = \frac{(1 - \cos(2\pi f))^2 + \sin^2(2\pi f)}{(2\pi f)^2} \cdot \frac{1 - \cos(\pi f)}{2(\pi f)^2} = \frac{(1 - \cos(2\pi f))(1 - \cos(\pi f))}{4(\pi f)^4}$$

$$\text{Answer: } S_Y(f) = \begin{cases} \frac{(1 - \cos(2\pi f))(1 - \cos(\pi f))}{4(\pi f)^4} & \text{om } f \neq 0 \\ 1/4 & \text{om } f = 0 \end{cases}$$

□