

# SOLUTIONS TO EXAM FOR STOCHASTIC MODELS IN DISCRETE TIME 3.75 ECTS

Master's program of Financial Mathematics  
August 11, 2008, 9.00 – 13.00

**Max number of points:** 30.

**Halmstad University grading bounds:** 12p  $\Rightarrow$  grade 3, 18p  $\Rightarrow$  grade 4, 24p  $\Rightarrow$  grade 5.

**ECTS bounds:** 12p  $\Rightarrow$  grade E, 15p  $\Rightarrow$  grade D, 18p  $\Rightarrow$  grade C, 21p  $\Rightarrow$  grade B, 24p  $\Rightarrow$  grade A.

**Allowed aids:** Summary of formulae attached to the exam, calculator and dictionary.

**Examiner:** Eric Järpe (035-16 76 53, 0702-822 844).

1. Prove the necessary part of the martingale criterion of absence of arbitrage. (6p)

**Solution:** (See p 413, *Essentials of Stochastic Finance. Facts, Models, Theory.* by A.N. Shiryaev.)  $\square$

2. Calculate the covariance function  $R(h) = \text{Cov}(X_t, X_{t+h})$  of an  $MA(1)$  process with parameters  $b_0 = b_1 = \sigma_\epsilon^2 = 1$ . (6p)

**Solution:**  $X_t = b_0 + b_1\epsilon_{t-1} + \sigma_\epsilon\epsilon_t = 1 + \epsilon_{t-1} + \epsilon_t$  where  $\{\epsilon_t\}$  is white noise.

$$D(X_t) = D(1 + \epsilon_{t-1} + \epsilon_t) = 0 + D(\epsilon_{t-1}) + D(\epsilon_t) = 2.$$

$$C(X_t, X_{t+1}) = C(1 + \epsilon_{t-1} + \epsilon_t, 1 + \epsilon_t + \epsilon_{t+1}) = C(\epsilon_t, \epsilon_t) = 1.$$

$$C(X_t, X_{t+h}) = C(1 + \epsilon_{t-1} + \epsilon_t, 1 + \epsilon_{t+h-1} + \epsilon_{t+h}) = 0 \text{ when } h \geq 2.$$

Since the covariance function is even we have that

$$R(h) = \begin{cases} 2 & \text{if } h = 0 \\ 1 & \text{if } |h| = 1 \\ 0 & \text{if } |h| \geq 2 \end{cases}$$

$\square$

3. Assume that stock market log returns,  $\{h_t\}$ , are distributed according to the  $HARCH(2)$  model with positive coefficients  $a_0, a_1, a_2$ . Show that

$$a_1 + 2a_2 < 1 \quad (6p)$$

**Solution:** In the  $HARCH(2)$  model  $h_n = \sigma_n\epsilon_n$  where  $\{\epsilon_n\}$  is white noise and  $\sigma_n^2 = a_0 + a_1h_{n-1}^2 + a_2(h_{n-1} + h_{n-2})^2 = a_0 + (a_1 + a_2)h_{n-1}^2 + a_2h_{n-2}^2 + 2a_2h_{n-1}h_{n-2}$ . We have that  $E(h_{n-1}h_{n-2}) = E(\sigma_{n-1}\epsilon_{n-1}h_{n-2}) = E(\epsilon_{n-1})E(\sigma_{n-1}h_{n-2}) = 0$  so  $E(h_n^2) = E(\sigma_n^2)E(\epsilon_n^2) = E(a_0 + (a_1 + a_2)h_{n-1}^2 + a_2h_{n-2}^2 + 2a_2h_{n-1}h_{n-2}) = a_0 + (a_1 + a_2)E(h_{n-1}^2) + a_2E(h_{n-2}^2) + 2a_2E(h_{n-1}h_{n-2}) = a_0 + (a_1 + 2a_2)E(h_n^2)$ , since  $E(h_{n-2}^2) = E(h_{n-1}^2) = E(h_n^2)$ . This means that  $E(h_n^2) = \frac{a_0}{1 - a_1 - 2a_2} > 0$  where  $a_0 > 0$  and thus also  $1 - a_1 - 2a_2 > 0$ , i.e.  $1 > a_1 + 2a_2$ .  $\square$

4. Assume  $\{X_n\}$  is a time homogeneous Markov chain with state space  $\{0, 1\}$  and transition probabilities  $P(X_{n+1} = 0 | X_n = 0) = 1$  and  $P(X_{n+1} = 0 | X_n = 1) = 0.01$ . Show that  $\{X_n\}$  is a submartingale. (6p)

**Solution:**  $P(X_{n+1} = 0 | X_n = 0) = 1 \Rightarrow P(X_{n+1} = 1 | X_n = 0) = 0$  and  $P(X_{n+1} = 0 | X_n = 1) = 0.01 \Rightarrow P(X_{n+1} = 1 | X_n = 1) = 0.99$ . We get  $E(X_{n+1} | X_n = 0) = 0 \cdot P(X_{n+1} = 0 | X_n = 0) + 1 \cdot P(X_{n+1} = 1 | X_n = 0) = 0 = X_n$  and  $E(X_{n+1} | X_n = 1) = 0 \cdot P(X_{n+1} = 0 | X_n = 1) + 1 \cdot P(X_{n+1} = 1 | X_n = 1) = 0.99 \leq 1 = X_n$ . Thus  $E(X_{n+1} | X_n) \leq X_n$ , i.e.  $\{X_n\}$  is a submartingale.  $\square$

5. Calculate  $E(X_n)$  where  $\{X_n : n \in \mathbb{N}\}$  is a geometric random walk<sup>1</sup>. (6p)

**Solution:**  $X_n = e^{R_n}$  where  $R_n = \sum_{k=1}^n U_k$  and  $U_k = \begin{cases} -1 & \text{w.p. } 1/2 \\ 1 & \text{w.p. } 1/2 \end{cases}$ .

Then, writing  $R_n$  as the linear combination  $2B_n - n$  of a random variable  $B_n \in \text{Bin}(n, \frac{1}{2})$  and using the binomial theorem, we have that

$$\begin{aligned}
 E(X_n) &= E(e^{R_n}) \\
 &= \sum_{j=-n, -n+2, \dots, n-2, n} e^j P(R_n = j) \\
 &= \sum_{j=-n+2, \dots, n-2, n} e^j P(2B_n - n = j) \\
 &= \sum_{j=-n, -n+2, \dots, n-2, n} e^j P(B_n = \frac{1}{2}(j+n)) \\
 &= \sum_{j=0}^n e^{2j-n} P(B_n = j) \\
 &= \sum_{j=0}^n e^{2j-n} \binom{n}{j} \left(\frac{1}{2}\right)^j \left(\frac{1}{2}\right)^{n-j} \\
 &= \sum_{j=0}^n (e^2)^j e^{-n} \binom{n}{j} \frac{1}{2^n} \\
 &= (2e)^{-n} \sum_{j=0}^n e^{-n} \binom{n}{j} (e^2)^j \cdot 1^{n-j} \\
 &= (2e)^{-n} (e^2 + 1)^n \\
 &= \left(\frac{e^2 + 1}{2e}\right)^n.
 \end{aligned}$$

Alternatively,  $E(X_n) = E(e^{\sum_{k=1}^n U_k}) = (E(e^{U_1}))^n = (e^{-1} \cdot \frac{1}{2} + e^1 \cdot \frac{1}{2})^n = \left(\frac{1+e^2}{2e}\right)^n$  (which is another way of writing  $\cosh^n(1)$ ).  $\square$

<sup>1</sup>A *geometric random walk* is a random process,  $\{X_n\}$ , satisfying  $X_n = e^{R_n}$  where the process  $\{R_n\}$  is a random walk