

# SOLUTIONS TO EXAM FOR STOCHASTIC MODELS IN DISCRETE TIME 3.75 ECTS

Master's program of Financial Mathematics

October 24, 2007, 9.00 – 13.00

**Max number of points:** 30.

**Halmstad University grading bounds:** 12p  $\Rightarrow$  grade 3, 18p  $\Rightarrow$  grade 4, 24p  $\Rightarrow$  grade 5.

**ECTS bounds:** 12p  $\Rightarrow$  grade E, 15p  $\Rightarrow$  grade D, 18p  $\Rightarrow$  grade C, 21p  $\Rightarrow$  grade B, 24p  $\Rightarrow$  grade A.

**Allowed aids:** Summary of formulae attached to the exam, calculator and dictionary.

**Examiner:** Eric Järpe (035-16 76 53, 0702-822 844).

1. Show that a local martingale is a generalised martingale, i.e. if  $X = \{X_n, \mathcal{F}_n : n \in \mathbb{Z}^+\}$  is a random process with  $E(|X_0|) < \infty$ , then you should prove that  $X \in \mathcal{M}_{loc} \Rightarrow X \in \mathcal{GM}$ , (where  $\mathcal{M}_{loc}$  is the class of local martingales and  $\mathcal{GM}$  is the class of generalised martingales). (4p)

**Solution:** (See p 98, *Essentials of Stochastic Finance. Facts, Models, Theory.* by A.N. Shiryaev.)  $\square$

2. Show that the variables of a process of the GARCH family are uncorrelated. (3p)

**Solution:** In the processes  $\{h_n\}$  of the GARCH family  $h_n = \sigma_n \epsilon_n$  for all  $n$  where  $\epsilon_n$  is white noise and  $m \leq n \Rightarrow \sigma_m \perp \epsilon_n$ . Thus  $E(h_n) = E(\sigma_n)E(\epsilon_n) = E(\sigma_n) \cdot 0 = 0$ . Now assume  $m < n$ . Then  $E(h_m h_n) = E(h_m \sigma_n \epsilon_n) = E(h_m \sigma_n)E(\epsilon_n) = E(h_m \sigma_n) \cdot 0 = 0$ . Since  $m \neq n \Rightarrow$  either  $m < n$  or  $m > n$  we have that  $E(h_m h_n) = 0$  for all  $m \neq n$  so  $Cov(h_m, h_n) = 0$  for all  $m \neq n$ , which is to say that the variables  $\{h_n\}$  are uncorrelated.  $\square$

3. Suppose  $\{h_t : t \in \mathbb{Z}^+\}$  is a random walk with  $h_0 = 0$  and

$$h_{t+1} = h_t + \begin{cases} 1 & \text{with probability } \frac{1}{2} \\ -2 & \text{with probability } \frac{1}{2} \end{cases}$$

(a) Calculate the covariance function  $R(s, t) = Cov(h_s, h_t)$ . (4p)

(b) Determine the deterministic function  $C(t)$  such that the process  $\{h_t + C(t) : t \in \mathbb{Z}^+\}$  is a martingale with respect to the flow  $\{\mathcal{F}_t\}$  where  $\mathcal{F}_t = \sigma(h_0, h_1, \dots, h_t)$ . (3p)

**Solution:**

(a) Let  $\{Z_t\}$  be a sequence of independent variables with  $Z_t = \begin{cases} 1 & \text{w.p. } 1/2 \\ -2 & \text{w.p. } 1/2 \end{cases}$ . Then  $h_t = h_{t-1} + Z_t = h_{t-2} + Z_{t-1} + Z_t = \dots = h_s + Z_{s+1} + \dots + Z_t$  and since  $\{Z_t\}$  are iid we have that  $h_t - h_s = Z_{s+1} + \dots + Z_t \perp Z_1 + \dots + Z_s = h_s$ , so assuming  $s < t$  we have that  $R(s, t) = \text{Cov}(h_s, h_t) = \text{Cov}(h_s, h_t - h_s + h_s) = \text{Cov}(h_s, h_t - h_s) + \text{Cov}(h_s, h_s) = 0 + D(h_s)$ . In the same way, for  $s \geq t$  we have  $R(s, t) = \dots = 0 + D(h_t)$ . This means that in general  $R(s, t) = D(h_{\min(s,t)})$ . Now,  $D(h_s) = D(Z_1 + \dots + Z_s) \stackrel{iid}{=} s D(Z_1)$  where  $D(Z_1) = E(Z_1^2) - E(Z_1)^2 = 1^2 \cdot \frac{1}{2} + (-2)^2 \cdot \frac{1}{2} - (1 \cdot \frac{1}{2} + (-2) \cdot \frac{1}{2})^2 = \frac{9}{4}$  so  $D(h_s) = \frac{9}{4}s$  and  $R(s, t) = D(h_{\min(s,t)}) = \frac{9}{4} \min(s, t)$ .

(b) We shall find  $C(t)$  such that  $\{h_t + C(t)\}$  is a martingale.  $E(h_{t+1} | \mathcal{F}_t) = E(h_t + Z_{t+1} | \mathcal{F}_t) = h_t + E(Z_{t+1}) = h_t + 1 \cdot \frac{1}{2} + (-2) \cdot \frac{1}{2} = h_t + \frac{1}{2}$ . Now, let  $g_t = h_t + C(t)$ . Then  $E(g_{t+1} | \mathcal{F}_t) = E(h_t + Z_{t+1} + C(t+1) | \mathcal{F}_t) = h_t + E(Z_{t+1}) + C(t+1) = h_t + \frac{1}{2} + C(t+1)$ . This equals  $g_t = h_t + C(t)$  if  $-\frac{1}{2} + C(t+1) = C(t)$ , i.e. if  $C(t+1) - C(t) = \frac{1}{2}$  for all  $t$ . To deduce that  $C(t) = \frac{1}{2}t + a$ , where  $a \in \mathbb{R}$  one can reason as follows.  
Alternative I: Differentiate w.r.t.  $t$  both sides of  $C(t+1) - C(t) = \frac{1}{2}$  (\*). This yields  $C'(t+1) = C'(t)$  for all  $t \Rightarrow C'(t) = b$  for some  $b \in \mathbb{R} \Rightarrow C(t) = a + bt$  where  $a, b \in \mathbb{R}$ . Now (\*)  $\Rightarrow b = \frac{1}{2}$ , so  $C(t) = a + \frac{1}{2}t$ .  
Alternative II: Since we are dealing with discrete time one may use the difference equation approach.  $C(t) = C(t-1) + \frac{1}{2} = C(t-2) + \frac{1}{2} + \frac{1}{2} = \dots = C(0) + t \cdot \frac{1}{2}$ . Thus  $C(t) = a + \frac{1}{2}t$  where  $a = C(0)$  is some real number.  $\square$

4. Let  $\{X_t : t \in \mathbb{Z}^+\}$  be an  $AR(1)$  process with white noise variance  $\sigma_\epsilon^2$  and with parameters  $a_0 = 0$  and  $a_1 = a$ . Assuming  $\{X_t\}$  is stationary, calculate

(a) the second moment of  $X_t$ , (4p)

(b) the fourth moment of  $X_t$ , (5p)

(c) the white noise variance  $\sigma_\epsilon^2$  if  $a = 0.9$  and  $D(X_t) = 1$ . (3p)

**Solution:**

(a)  $X_t = aX_{t-1} + \sigma_\epsilon \epsilon_t$  where  $\{\epsilon_t\}$  are independent and  $\epsilon_t \in N(0, 1)$ .  
 $m_1 = E(X_t) = aE(X_{t-1}) + \sigma_\epsilon E(\epsilon_t) \stackrel{stationarity}{\Rightarrow} m_1(1 - a) = 0 \Rightarrow E(X_t) = m_1 = 0$ .  
 $m_2 = E(X_t^2) = E(a^2 X_{t-1}^2 + 2a\sigma_\epsilon X_{t-1}\epsilon_t + \sigma_\epsilon^2 \epsilon_t^2) = a^2 E(X_{t-1}^2) + 2a\sigma_\epsilon E(X_{t-1})E(\epsilon_t) + \sigma_\epsilon^2 E(\epsilon_t^2) \stackrel{stationarity}{\Rightarrow} m_2(1 - a^2) = 0 + \sigma_\epsilon^2 \cdot 1 \Rightarrow E(X_t^2) = m_2 = \frac{\sigma_\epsilon^2}{1 - a^2}$ .

$$\begin{aligned}
\text{(b) } m_4 &= E(X_t^4) = E((a^2 X_{t-1}^2 + 2a\sigma_\epsilon X_{t-1}\epsilon_t + \sigma_\epsilon^2 \epsilon_t^2)^2) = \\
&E(a^4 X_{t-1}^4 + 4a^2 \sigma_\epsilon^2 X_{t-1}^2 \epsilon_t^2 + \sigma_\epsilon^4 \epsilon_t^4 + 4a^3 \sigma_\epsilon X_{t-1}^3 \epsilon_t + 2a^2 \sigma_\epsilon^2 X_{t-1}^2 \epsilon_t^2 + 4a\sigma_\epsilon^3 X_{t-1} \epsilon_t^3) = \\
&= a^4 E(X_{t-1}^4) + 4a^2 \sigma_\epsilon^2 E(X_{t-1}^2) E(\epsilon_t^2) + \sigma_\epsilon^4 E(\epsilon_t^4) + 4a^3 \sigma_\epsilon E(X_{t-1}^3) E(\epsilon_t) + \\
&+ 2a^2 \sigma_\epsilon^2 E(X_{t-1}^2) E(\epsilon_t^2) + 4a\sigma_\epsilon^3 E(X_{t-1}) E(\epsilon_t^3) = \\
&= a^4 m_4 + 4a^2 \sigma_\epsilon^2 \frac{\sigma_\epsilon^2}{1-a^2} + 3\sigma_\epsilon^4 + 0 + 2a^2 \sigma_\epsilon^2 \frac{\sigma_\epsilon^2}{1-a^2} + 0 \\
&\text{(since } \{X_t\} \text{ is stationary and } E(\epsilon_t) = E(X_t) = 0) \\
&\Rightarrow m_4(1-a^4) = \frac{6a^2 \sigma_\epsilon^4}{1-a^2} + 3\sigma_\epsilon^4 \\
&\Rightarrow E(X_t^4) = m_4 = \\
&= \frac{6a^2 \sigma_\epsilon^4 + 3\sigma_\epsilon^4(1-a^2)}{(1-a^2)(1-a^4)} = \frac{3\sigma_\epsilon^4(1+a^2)}{(1-a^2)(1-a^2)(1+a^2)} = \frac{3\sigma_\epsilon^4}{(1-a^2)^2}.
\end{aligned}$$

$$\begin{aligned}
\text{(c) } X_t &= aX_{t-1} + \sigma_\epsilon \epsilon_t \\
1 &= D(X_t) = D(aX_{t-1} + \sigma_\epsilon \epsilon_t) = a^2 D(X_{t-1}) + \sigma_\epsilon^2 D(\epsilon_t) = a^2 + \sigma_\epsilon^2. \quad a = 0.9 \Rightarrow \\
\sigma_\epsilon^2 &= 1 - 0.9^2 = 0.19. \quad \square
\end{aligned}$$

5. Consider a  $(B, S)$ -market model  $(\{B_n\}, \{S_n\})$  where  $B_0 > 0$ ,  $S_0 > 0$ ,  $B_{n+1} = (1+r)B_n$ ,  $S_{n+1} = (1+\rho)S_n$  for all  $n$  and  $r > -1$ ,  $\rho > -1$ . Show that  $E(\rho) \geq r$  implies that the process  $\{\frac{S_n}{B_n} : n \in \mathbb{Z}^+\}$  is a submartingale with respect to the flow  $\mathcal{F}_n = \sigma(\frac{S_0}{B_0}, \frac{S_1}{B_1}, \dots, \frac{S_n}{B_n})$ . (4p)

**Solution:** Assume  $E(\rho) \geq r$ . Then

$$\begin{aligned}
E\left(\frac{S_{n+1}}{B_{n+1}} \mid \mathcal{F}_n\right) &= E\left(\frac{(1+\rho)S_n}{(1+r)B_n} \mid \mathcal{F}_n\right) \\
&= \frac{1+E(\rho)}{1+r} \cdot \frac{S_n}{B_n} \\
&\geq \frac{1+r}{1+r} \cdot \frac{S_n}{B_n} \\
&= \frac{S_n}{B_n}
\end{aligned}$$

$\Rightarrow \{\frac{S_n}{B_n}\}$  is a submartingale. □