

# EXAM FOR STOCHASTIC MODELS IN DISCRETE TIME 3.75 ECTS

Master's program of Financial Mathematics  
January 4, 2008, 9.00 – 13.00

**Max number of points:** 30.

**Halmstad University grading bounds:** 12p  $\Rightarrow$  grade 3, 18p  $\Rightarrow$  grade 4, 24p  $\Rightarrow$  grade 5.

**ECTS bounds:** 12p  $\Rightarrow$  grade E, 15p  $\Rightarrow$  grade D, 18p  $\Rightarrow$  grade C, 21p  $\Rightarrow$  grade B, 24p  $\Rightarrow$  grade A.

**Allowed aids:** Summary of formulae attached to the exam, calculator and dictionary.

**Examiner:** Eric Järpe (035-16 76 53, 0702-822 844).

For each problem a *complete* solution should be given. All solutions should be thoroughly presented.

Each solution should start at the top of a new sheet of paper. Only one solution a sheet.

The proper solutions will be available on the internet at <http://www.hh.se/staff/erja>  $\rightarrow$  Teaching

$\rightarrow$  Financial Mathematics  $\rightarrow$  Stochastic models  $\rightarrow$  Previous exams  $\rightarrow$  080104: Solution

1. Prove that a generalised martingale is a martingale transformation. (5p)

2. Let  $\{N_t : t \in \mathbb{R}^+\}$  be a Poisson process and define  $X_t = N_t - a(t)$  for all  $t \in \mathbb{Z}^+$ . Determine the function  $a(t)$  such that  $\{X_t\}$  is a martingale with respect to the flow  $\mathcal{F}_t = \sigma(N_t, N_{t-1}, \dots, N_1)$ . (4p)

3. Calculate  $D(X_t)$  if  $\{X_t\}$  is an  $AR(1)$  process with  $a_0 = a_1 = \sigma_\epsilon^2 = \frac{1}{2}$ . (5p)

4. Let  $\{h_t : t \in \mathbb{Z}\}$  be an  $ARCH(1)$  process.

(a) Prove that  $h_t$  and  $h_{t-1}$  are uncorrelated. (3p)

(b) Calculate  $E(h_t^2 h_{t-1}^2)$ . (6p)

5. Suppose the price,  $S_t$ , in dollars of an option are developing according to  $S_t = 100 + W_t$  where  $\{W_t\}$  is a standard Wiener process. Suppose also that you are given the following opportunity: at the time  $t = 0$  you are offered to buy the option for the price \$99 with the obligation that, at time  $t = 100$ , you lose  $S_t - 99$  if  $S_t < 99$  and you win  $a(S_t - 99)$  if  $S_t \geq 99$  where  $a \in \mathbb{R}$ . Determine how large  $a$  should be chosen so that there is no arbitrage for the buyer. (7p)

*GOOD LUCK!*



<b>Expected value and variance</b>	$X$ discrete:	The <b>expected value</b> of $X$ : $\mu = E(X) = \sum_{x \in \Omega} x p(x)$ . <b>Variance</b> of $X$ : $\sigma^2 = D(X) = \sum_{x \in \Omega} (x - \mu)^2 p(x)$ .
	$X$ continuous:	<b>Expected value</b> of $X$ : $\mu = E(X) = \int_{x \in \Omega} x f(x) dx$ . <b>Variance</b> of $X$ : $\sigma^2 = D(X) = \int_{x \in \Omega} (x - \mu)^2 f(x) dx$ .
		<b>Skewness</b> of $X$ : $\gamma_1 = S(X) = \frac{E((X-\mu)^3)}{E((X-\mu)^2)^{3/2}}$
		<b>Kurtosis</b> of $X$ : $\gamma_2 = K(X) = \frac{E((X-\mu)^4)}{E((X-\mu)^2)^2} - 3$ $\gamma_2 > 0 \Rightarrow X$ is <b>leptokurtic</b> (heavy tails) $\gamma_2 = 0 \Rightarrow X$ is <b>mesokurtic</b> $\gamma_2 < 0 \Rightarrow X$ is <b>platykurtic</b> (light tails)
		<b>Covariance</b> of $X$ and $Y$ : $C(X, Y) = E((X - \mu_x)(Y - \mu_y))$
		<b>Correlation</b> of $X$ and $Y$ : $\rho = \frac{C(X, Y)}{\sqrt{D(X)D(Y)}}$
<b>Standarddev.</b>		$\sigma = \sqrt{D(X)}$ .
<b>Linearity:</b>		$E(aX + bY) = a E(X) + b E(Y)$ for all random variables $X$ and $Y$ and real numbers $a$ and $b$ . If $X, Y$ indep. then $D(aX + bY) = a^2 D(X) + b^2 D(Y)$ .
<b>Rules:</b>		$E(g(X)) = \int_{\mathbb{R}} g(x) f(x) dx$ $E(X A) = E(X \cdot I(A))$ $D(X) = E(X^2) - (E(X))^2$ $C(X, Y) = \int \int_{\mathbb{R}^2} xy f(x, y) - E(X)E(Y)$ $C(\sum_i a_i X_i, \sum_k b_k Y_k) = \sum_i \sum_k a_i b_k C(X_i, Y_k)$

**Normal distribution** denoted by  $N(\mu, \sigma)$  where  $\mu$  is the expected value and  $\sigma$  is the standard deviation  $N(0, 1)$  is called **standard normal distribution** with density function  $\Phi(x)$

If  $X \in N(\mu, \sigma)$  then  $P(X \leq x) = \Phi\left(\frac{x-\mu}{\sigma}\right)$

*Symmetry:*  $\Phi(-x) = 1 - \Phi(x)$

*Probabilities:*  $P(a \leq X \leq b) = \Phi\left(\frac{b-\mu}{\sigma}\right) - \Phi\left(\frac{a-\mu}{\sigma}\right)$

**Def** The random variables  $X_1, X_2, \dots, X_n$  are a **sample** of  $X$  if all variables,  $X_i$ , is distributed as  $X$ ,  $i = 1, \dots, n$ , and all variables are independent of each other at all levels.

**CLT** *Central Limit Theorem*

If  $X_1, \dots, X_n$  is a sample where  
 $E(X_i) = \mu$  and  $D(X_i) = \sigma^2$ ,  $i = 1, \dots, n$

then  $P\left(\frac{\sqrt{n}}{\sigma}(\bar{X} - \mu) \leq x\right) \rightarrow \Phi(x)$  as  $n \rightarrow \infty$ .

This implies that  $\sum_{i=1}^n X_i$  is approximately  $N(n\mu, \sqrt{n}\sigma)$  and  $\bar{X}$  is approximately  $N(\mu, \sigma/\sqrt{n})$  for large  $n$ .

**Thm** If  $\{X_t\}$  has independent, stationary increments and  $X_0 = 0$   
then  $R_X(s, t) = \min(s, t) \cdot D(X_1)$ .

**Thm** *Fatou lemma*

$$X_n \geq 0 \Rightarrow E(\liminf X_n) \leq \liminf E(X_n)$$

**Thm** *Dominated-convergence theorem*

$$\text{If } \exists C < \infty : |X_n| < C \text{ for all } n \text{ and } \exists \lim_{n \rightarrow \infty} X_n, \text{ then } \lim_{n \rightarrow \infty} E(X_n) = E(\lim_{n \rightarrow \infty} X_n).$$

**Def** **White noise** in discrete time is a sequence  $\{\epsilon_n\}_{n=-\infty}^{\infty}$  of

independent random variables,  $\epsilon_n$ , such that  $E(\epsilon_n) = 0$  and  $D(\epsilon_n) = \sigma_\epsilon^2$ .

Let the **lag operator**  $L$  be defined by  $L^k(X_n) = X_{n-k}$  and let  $\alpha$  and  $\beta$  be the polynomials  $\alpha(z) = 1 - \sum_{k=1}^p a_k z^k$  and  $\beta(z) = 1 + \sum_{k=1}^q b_k z^k$ .

**Def** Let  $\{X_n\} = \{X_n : n \in \mathbb{Z}\}$  be a sequence of variables and let  $\{\epsilon_n\}$  be white noise such that  $\epsilon_m \perp X_n$  whenever  $m > n$ . Then  $\{X_n\}$  is an

**AR(p) process** if  $\alpha(L)X_n = a_0 + \epsilon_n$

**MA(q) process** if  $X_n = \beta(L)\epsilon_n$

**ARMA(p, q) process** if  $\alpha(L)X_n = \beta(L)\epsilon_n$

**ARIMA(p, d, q) process** if  $\{(1-L)^d X_n : n \in \mathbb{Z}\}$  is an  $ARMA(p, q)$  process for all  $n \in \mathbb{Z}$ .

**Thm** *Yule Walker equations*

For an  $AR(p)$  process the covariance function is

$$R(k) - \sum_{j=1}^p a_j R(k-j) = \begin{cases} \sigma_\epsilon^2 & \text{for } k = 0 \\ 0 & \text{for } k = 1, 2, \dots \end{cases}$$

**Thm** For an  $MA(q)$  process the covariance function is

$$R(k) = \begin{cases} \sigma_\epsilon^2 \sum_{i-j=k} c_i c_j & \text{for } |k| \leq q \\ 0 & \text{for } |k| > q \end{cases}$$

**Def** Let  $\{X_n\} = \{X_n : n \in \mathbb{Z}\}$  be a sequence of variables satisfying  $X_n = \sigma_n \epsilon_n$  for all  $n$  and where  $\{\epsilon_n\}$  is white noise such that  $\epsilon_m \perp X_n$  whenever  $m > n$ . Then  $\{X_n\}$  is

an **ARCH(p) process** if  $\sigma_n^2 = a_0 + \sum_{k=1}^p a_k X_{n-k}^2$

a **GARCH(p, q) process** if  $\sigma_n^2 = a_0 + \sum_{k=1}^p a_k X_{n-k}^2 + \sum_{k=1}^q b_k \sigma_{n-k}^2$

a **HARCH(p) process** if  $\sigma_n^2 = a_0 + \sum_{k=1}^p a_k (\sum_{j=1}^k X_{n-j})^2$

a **Stochastic volatility process of order p** if  $\sigma_n^2 = e^{\Delta_n}$  and

$\Delta_n = a_0 + \sum_{k=1}^p a_k \Delta_{n-k} + c \delta_n$  where  $\{\delta_n\}$  is white noise independent of  $\{\epsilon_n\}$  for all  $n \in \mathbb{Z}$ .

## Trigonometrics

$$\cos^2 \alpha + \sin^2 \alpha = 1$$

$$\cos(\alpha + \beta) = \cos \alpha \cos \beta - \sin \alpha \sin \beta \quad \cos \alpha = \frac{1}{2}(e^{i\alpha} + e^{-i\alpha})$$

$$\sin(\alpha + \beta) = \sin \alpha \cos \beta + \cos \alpha \sin \beta \quad \sin \alpha = \frac{1}{2i}(e^{i\alpha} - e^{-i\alpha})$$

## Inference

**Thm** (Estimation of expected value)

$$\begin{aligned} \text{If } \hat{m}_n &= \frac{1}{n} \sum_{t=1}^n X_t \\ \text{then } D(\hat{m}_n) &= \frac{1}{n^2} \sum_{k=-n+1}^{n-1} (n - |k|) R_X(k) \\ nD(\hat{m}_n) &\approx \sum_{k=-\infty}^{\infty} R_X(k) \text{ for large } n \end{aligned}$$

**Def** If  $\mathbf{X} = (X_1, \dots, X_n)$  is a sample of the variable  $X$  distributed according to the density function  $f_X(x; \theta)$ , then the **likelihood function** of  $\theta$  is the joint density function of  $\mathbf{X}$ ,  $L(\theta) = \prod_i f(x_i; \theta)$ , as a function of the parameter  $\theta$ . The value of  $\theta$  which maximises the likelihood (or equivalently the log likelihood  $\ell(\theta) = \ln L(\theta)$ ) is the **maximum likelihood estimator (MLE)**  $\hat{\theta}$  of  $\theta$ .

**Def** A point estimator,  $\theta^*$ , of a parameter  $\theta$  is **unbiased** if  $E(\theta^*) = \theta$ . If  $\theta_1^*$  and  $\theta_2^*$  are unbiased estimators of  $\theta$ , then  $\theta_1^*$  is **better/more efficient** than  $\theta_2^*$  om  $D(\theta_1^*) < D(\theta_2^*)$ .

## Distributions, expected values and variances

	$X$	$p(x), f(x)$	$E(X)$	$D(X)$
Discrete distributions	Unif( $N$ )	$1/N$ $x = 1, 2, \dots, N$	$(N + 1)/2$	$(N^2 - 1)/12$
	Bin( $n, p$ )	$\binom{n}{x} p^x (1 - p)^{n-x}$ $x = 0, 1, 2, \dots, n$	$np$	$np(1 - p)$
	Poi( $\lambda$ )	$e^{-\lambda} \lambda^x / x!$ $x = 0, 1, 2, \dots$	$\lambda$	$\lambda$
	Geo( $\pi$ )	$(1 - \pi)^{x-1} \pi$ $x = 1, 2, 3, \dots$	$1/\pi$	$(1 - \pi)/\pi^2$
Cont. distributions	R( $a, b$ )	$1/(b - a)$ $a \leq x \leq b$	$(a + b)/2$	$(a - b)^2/12$
	Exp( $\lambda$ )	$\lambda e^{-\lambda x}$ $x > 0$	$1/\lambda$	$1/\lambda^2$
	N( $\mu, \sigma$ )	$(\sigma\sqrt{2\pi})^{-1} e^{-(x-\mu)^2/(2\sigma^2)}$ $x \in \mathbb{R}$	$\mu$	$\sigma^2$

# Normal distribution values

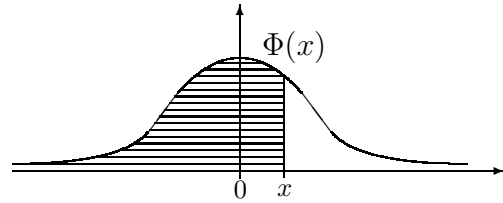


Table over values of  $\Phi(x) = P(X \leq x)$  where  $X \in N(0, 1)$ . For  $x < 0$ , use the relation  $\Phi(x) = 1 - \Phi(-x)$ .

$x$	+0.00	+0.01	+0.02	+0.03	+0.04	+0.05	+0.06	+0.07	+0.08	+0.09
0.0	0.5000	0.5040	0.5080	0.5120	0.5160	0.5199	0.5239	0.5279	0.5319	0.5359
0.1	0.5398	0.5438	0.5478	0.5517	0.5557	0.5596	0.5636	0.5675	0.5714	0.5753
0.2	0.5793	0.5832	0.5871	0.5910	0.5948	0.5987	0.6026	0.6064	0.6103	0.6141
0.3	0.6179	0.6217	0.6255	0.6293	0.6331	0.6368	0.6406	0.6443	0.6480	0.6517
0.4	0.6554	0.6591	0.6628	0.6664	0.6700	0.6736	0.6772	0.6808	0.6844	0.6879
0.5	0.6915	0.6950	0.6985	0.7019	0.7054	0.7088	0.7123	0.7157	0.7190	0.7224
0.6	0.7257	0.7291	0.7324	0.7357	0.7389	0.7422	0.7454	0.7486	0.7517	0.7549
0.7	0.7580	0.7611	0.7642	0.7673	0.7704	0.7734	0.7764	0.7794	0.7823	0.7852
0.8	0.7881	0.7910	0.7939	0.7967	0.7995	0.8023	0.8051	0.8078	0.8106	0.8133
0.9	0.8159	0.8186	0.8212	0.8238	0.8264	0.8289	0.8315	0.8340	0.8365	0.8389
1.0	0.8413	0.8438	0.8461	0.8485	0.8508	0.8531	0.8554	0.8577	0.8599	0.8621
1.1	0.8643	0.8665	0.8686	0.8708	0.8729	0.8749	0.8770	0.8790	0.8810	0.8830
1.2	0.8849	0.8869	0.8888	0.8907	0.8925	0.8944	0.8962	0.8980	0.8997	0.9015
1.3	0.9032	0.9049	0.9066	0.9082	0.9099	0.9115	0.9131	0.9147	0.9162	0.9177
1.4	0.9192	0.9207	0.9222	0.9236	0.9251	0.9265	0.9279	0.9292	0.9306	0.9319
1.5	0.9332	0.9345	0.9357	0.9370	0.9382	0.9394	0.9406	0.9418	0.9429	0.9441
1.6	0.9452	0.9463	0.9474	0.9484	0.9495	0.9505	0.9515	0.9525	0.9535	0.9545
1.7	0.9554	0.9564	0.9573	0.9582	0.9591	0.9599	0.9608	0.9616	0.9625	0.9633
1.8	0.9641	0.9649	0.9656	0.9664	0.9671	0.9678	0.9686	0.9693	0.9699	0.9706
1.9	0.9713	0.9719	0.9726	0.9732	0.9738	0.9744	0.9750	0.9756	0.9761	0.9767
2.0	0.9772	0.9778	0.9783	0.9788	0.9793	0.9798	0.9803	0.9808	0.9812	0.9817
2.1	0.9821	0.9826	0.9830	0.9834	0.9838	0.9842	0.9846	0.9850	0.9854	0.9857
2.2	0.9861	0.9864	0.9868	0.9871	0.9875	0.9878	0.9881	0.9884	0.9887	0.9890
2.3	0.9893	0.9896	0.9898	0.9901	0.9904	0.9906	0.9909	0.9911	0.9913	0.9916
2.4	0.9918	0.9920	0.9922	0.9925	0.9927	0.9929	0.9931	0.9932	0.9934	0.9936
2.5	0.9938	0.9940	0.9941	0.9943	0.9945	0.9946	0.9948	0.9949	0.9951	0.9952
2.6	0.9953	0.9955	0.9956	0.9957	0.9959	0.9960	0.9961	0.9962	0.9963	0.9964
2.7	0.9965	0.9966	0.9967	0.9968	0.9969	0.9970	0.9971	0.9972	0.9973	0.9974
2.8	0.9974	0.9975	0.9976	0.9977	0.9977	0.9978	0.9979	0.9979	0.9980	0.9981
2.9	0.9981	0.9982	0.9982	0.9983	0.9984	0.9984	0.9985	0.9985	0.9986	0.9986

$x$	+0.0	+0.1	+0.2	+0.3	+0.4	+0.5	+0.6	+0.7	+0.8	+0.9
3	0.9987	0.9990	0.9993	0.9995	0.9997	0.9998	0.9998	0.9999	0.9999	1.0000

## Percentiles:

Some values of  $\lambda_\alpha$  such that  $P(X > \lambda_\alpha) = \alpha$  where  $X \in N(0, 1)$

$\alpha$	$\lambda_\alpha$	$\alpha$	$\lambda_\alpha$
0.1	1.281552	0.005	2.575829
0.05	1.644854	0.001	3.090232
0.025	1.959964	0.0005	3.290527
0.01	2.326348	0.0001	3.719016

## $t$ percentiles

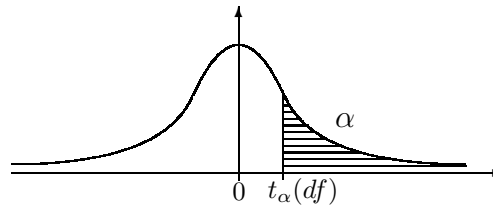


Table over values of  $t_\alpha(df)$ .

$df$	$\alpha$	0.25	0.10	0.05	0.025	0.02	0.01	0.005	0.001
1		1.0000	3.0777	6.3138	12.7062	15.8945	31.8205	63.6567	318.3088
2		0.8165	1.8856	2.9200	4.3027	4.8487	6.9646	9.9248	22.3271
3		0.7649	1.6377	2.3534	3.1824	3.4819	4.5407	5.8409	10.2145
4		0.7407	1.5332	2.1318	2.7764	2.9986	3.7470	4.6041	7.1732
5		0.7267	1.4759	2.0150	2.5706	2.7565	3.3649	4.0322	5.8934
6		0.7176	1.4398	1.9432	2.4469	2.6122	3.1427	3.7074	5.2076
7		0.7111	1.4149	1.8946	2.3646	2.5168	2.9980	3.4995	4.7853
8		0.7064	1.3968	1.8595	2.3060	2.4490	2.8965	3.3554	4.5008
9		0.7027	1.3830	1.8331	2.2622	2.3984	2.8214	3.2498	4.2968
10		0.6998	1.3722	1.8125	2.2281	2.3593	2.7638	3.1693	4.1437
12		0.6955	1.3562	1.7823	2.1788	2.3027	2.6810	3.0545	3.9296
14		0.6924	1.3450	1.7613	2.1448	2.2638	2.6245	2.9768	3.7874
17		0.6892	1.3334	1.7396	2.1098	2.2238	2.5669	2.8982	3.6458
20		0.6870	1.3253	1.7247	2.0860	2.1967	2.5280	2.8453	3.5518
25		0.6844	1.3163	1.7081	2.0595	2.1666	2.4851	2.7874	3.4502
30		0.6828	1.3104	1.6973	2.0423	2.1470	2.4573	2.7500	3.3852
50		0.6794	1.2987	1.6759	2.0086	2.1087	2.4033	2.6778	3.2614
100		0.6770	1.2901	1.6602	1.9840	2.0809	2.3642	2.6259	3.1737

## $\chi^2$ percentiles

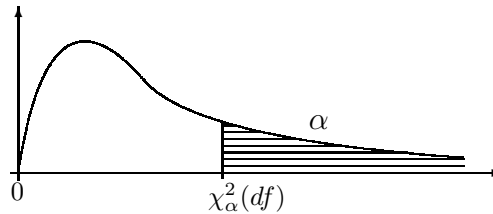


Table over values of  $\chi_\alpha^2(df)$ .

$df$	$\alpha$	0.999	0.995	0.99	0.95	0.05	0.01	0.005	0.001
1		0.0000	0.0000	0.0002	0.0039	3.8415	6.6349	7.8794	10.8276
2		0.0020	0.0100	0.0201	0.1026	5.9915	9.2103	10.5966	13.8155
3		0.0243	0.0717	0.1148	0.3518	7.8147	11.3449	12.8382	16.2662
4		0.0908	0.2070	0.2971	0.7107	9.4877	13.2767	14.8603	18.4668
5		0.2102	0.4117	0.5543	1.1455	11.0705	15.0863	16.7496	20.5150
6		0.3811	0.6757	0.8721	1.6354	12.5916	16.8119	18.5476	22.4577
7		0.5985	0.9893	1.2390	2.1673	14.0671	18.4753	20.2777	24.3219
8		0.8571	1.3444	1.6465	2.7326	15.5073	20.0902	21.9550	26.1245
9		1.1519	1.7349	2.0879	3.3251	16.9190	21.6660	23.5894	27.8772
10		1.4787	2.1559	2.5582	3.9403	18.3070	23.2093	25.1882	29.5883
12		2.2142	3.0738	3.5706	5.2260	21.0261	26.2170	28.2995	32.9095
14		3.0407	4.0747	4.6604	6.5706	23.6848	29.1412	31.3193	36.1233
17		4.4161	5.6972	6.4078	8.6718	27.5871	33.4087	35.7185	40.7902
20		5.9210	7.4338	8.2604	10.8508	31.4104	37.5662	39.9968	45.3147
25		8.6493	10.5197	11.5240	14.6114	37.6525	44.3141	46.9279	52.6197
30		11.5880	13.7867	14.9535	18.4927	43.7730	50.8922	53.6720	59.7031
50		24.6739	27.9907	29.7067	34.7643	67.5048	76.1539	79.4900	86.6608
100		61.9179	67.3276	70.0649	77.9295	124.342	135.807	140.169	149.449

# Values of the Poisson distribuion

Table over values of  $F(x) = P(X \leq x)$  where  $X \in Po(\lambda)$ .

$\lambda$	0	1	2	3	4	5	6	7	8	9	10	11	12	13
0.5	0.607	0.910	0.986	0.998	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000
1	0.368	0.736	0.920	0.981	0.996	0.999	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000
2	0.135	0.406	0.677	0.857	0.947	0.983	0.995	0.999	1.000	1.000	1.000	1.000	1.000	1.000
3	0.050	0.199	0.423	0.647	0.815	0.916	0.966	0.988	0.996	0.999	1.000	1.000	1.000	1.000
4	0.018	0.092	0.238	0.433	0.629	0.785	0.889	0.949	0.979	0.992	0.997	0.999	1.000	1.000
5	0.007	0.040	0.125	0.265	0.440	0.616	0.762	0.867	0.932	0.968	0.986	0.995	0.998	0.999
6	0.002	0.017	0.062	0.151	0.285	0.446	0.606	0.744	0.847	0.916	0.957	0.980	0.991	0.996

# Values of the Binomial distribution

Table over values of  $P(x) = P(X \leq x)$  where  $X \in Bin(n, p)$ .

For  $p > 0.5$ , use the relation  $P(X \leq x) = P(Y \geq n - x)$  where  $Y \in Bin(n, 1 - p)$ .

$n$	$p$	0	1	2	3	4	5	6	7	8	9	10
3	0.1	0.729	0.972	0.999	1.000	–	–	–	–	–	–	–
	0.2	0.512	0.896	0.992	1.000	–	–	–	–	–	–	–
	0.3	0.343	0.784	0.973	1.000	–	–	–	–	–	–	–
	0.4	0.216	0.648	0.936	1.000	–	–	–	–	–	–	–
	0.5	0.125	0.500	0.875	1.000	–	–	–	–	–	–	–
4	0.1	0.656	0.948	0.996	1.000	1.000	–	–	–	–	–	–
	0.2	0.410	0.819	0.973	0.998	1.000	–	–	–	–	–	–
	0.3	0.240	0.652	0.916	0.992	1.000	–	–	–	–	–	–
	0.4	0.130	0.475	0.821	0.974	1.000	–	–	–	–	–	–
	0.5	0.062	0.312	0.688	0.938	1.000	–	–	–	–	–	–
5	0.1	0.590	0.919	0.991	1.000	1.000	1.000	–	–	–	–	–
	0.2	0.328	0.737	0.942	0.993	1.000	1.000	–	–	–	–	–
	0.3	0.168	0.528	0.837	0.969	0.998	1.000	–	–	–	–	–
	0.4	0.078	0.337	0.683	0.913	0.990	1.000	–	–	–	–	–
	0.5	0.031	0.188	0.500	0.812	0.969	1.000	–	–	–	–	–
6	0.1	0.531	0.886	0.984	0.999	1.000	1.000	1.000	–	–	–	–
	0.2	0.262	0.655	0.901	0.983	0.998	1.000	1.000	–	–	–	–
	0.3	0.118	0.420	0.744	0.930	0.989	0.999	1.000	–	–	–	–
	0.4	0.047	0.233	0.544	0.821	0.959	0.996	1.000	–	–	–	–
	0.5	0.016	0.109	0.344	0.656	0.891	0.984	1.000	–	–	–	–
7	0.1	0.478	0.850	0.974	0.997	1.000	1.000	1.000	1.000	–	–	–
	0.2	0.210	0.577	0.852	0.967	0.995	1.000	1.000	1.000	–	–	–
	0.3	0.082	0.329	0.647	0.874	0.971	0.996	1.000	1.000	–	–	–
	0.4	0.028	0.159	0.420	0.710	0.904	0.981	0.998	1.000	–	–	–
	0.5	0.008	0.062	0.227	0.500	0.773	0.938	0.992	1.000	–	–	–
8	0.1	0.430	0.813	0.962	0.995	1.000	1.000	1.000	1.000	1.000	–	–
	0.2	0.168	0.503	0.797	0.944	0.990	0.999	1.000	1.000	1.000	–	–
	0.3	0.058	0.255	0.552	0.806	0.942	0.989	0.999	1.000	1.000	–	–
	0.4	0.017	0.106	0.315	0.594	0.826	0.950	0.991	0.999	1.000	–	–
	0.5	0.004	0.035	0.145	0.363	0.637	0.855	0.965	0.996	1.000	–	–
9	0.1	0.387	0.775	0.947	0.992	0.999	1.000	1.000	1.000	1.000	1.000	–
	0.2	0.134	0.436	0.738	0.914	0.980	0.997	1.000	1.000	1.000	1.000	–
	0.3	0.040	0.196	0.463	0.730	0.901	0.975	0.996	1.000	1.000	1.000	–
	0.4	0.010	0.071	0.232	0.483	0.733	0.901	0.975	0.996	1.000	1.000	–
	0.5	0.002	0.020	0.090	0.254	0.500	0.746	0.910	0.980	0.998	1.000	–
10	0.1	0.349	0.736	0.930	0.987	0.998	1.000	1.000	1.000	1.000	1.000	1.000
	0.2	0.107	0.376	0.678	0.879	0.967	0.994	0.999	1.000	1.000	1.000	1.000
	0.3	0.028	0.149	0.383	0.650	0.850	0.953	0.989	0.998	1.000	1.000	1.000
	0.4	0.006	0.046	0.167	0.382	0.633	0.834	0.945	0.988	0.998	1.000	1.000
	0.5	0.001	0.011	0.055	0.172	0.377	0.623	0.828	0.945	0.989	0.999	1.000